



A Technical Report Grant No. N00014-86-K-0742 September 1, 1986 - August 31, 1988

AN ALGORITHM FOR DETECTING CHANGES IN BATTLE SCENARIA

Submitted to:

Director National Research Laboratory Washington, D. C. 20375

Attention: Code 2627

Submitted by:

P. P. Kazakos Professor

R. K. Bansal Graduate Assistant



Report No. UVA/525415/EE88/109 October 1987

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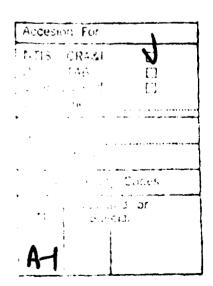
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#### I. Introduction

~ In distributed tactical decision making, commanders make decisions based on data generated by their local territory as well as on decisions and information communicated to them by other commanders. The timely aspect of the decisions is controlled by the rate with which the commanders receive data, which is in turn controlled by the deployed transmission algorithms. (such as those in [3] and [4]). The accuracy of the decisions, on the other hand, depends heavily on the way that the decision makers perceive their environment. This perception corresponds to a number of alternative models, where the latter are a priori developed based on the various battle scenaria. Each such model is associated with an appropriate decision mode. Furthermore, a model (or battle scenario) may shift to another such model, at a random time. In the event of such a shift, it is thus crucial that the commander be alerted, for adaptation to the appropriate decision mode. In this paper, we describe an algorithm which can be deployed in an automatic fashion; to alert the commanders about changes in battle scenaria.

#### II. The Algorithm

Each battle scenario is modelled by a stochastic process. Let the various battle scenaria be indexed, say from 1 to M. Then, the ith battle scenario is represented by a distribution  $F_i(x_1,...,x_n,...)$ , where  $x_1,...,x_n,...$  are the data outputs observed. Let us assume that the above distributions possess density functions, and let then  $f_i(x_n | x_1,...,x_{n-1})$  denote the density of the datum  $x_n$ , given the data  $x_1$ ...,  $x_{n-1}$ , and given that the ith battle scenario is present.

Let the jth battle scenario be known to be present at time zero. Our objective is to decide if a change in this scenario occurs, and if the answer is yes, then to decide what is the new scenario. Furthermore, if a change does occur, we wish to decide affirmatively as soon from the instant of its occurence as possible. Let us define,

$$g_{ij}(x_1,...,x_n) = \log \frac{f_i(x_n | x_1,...,x_{n-1})}{f_j(x_n | x_1,...,x_{n-1})}$$
(1)

T<sub>ij</sub><sup>(n)</sup> such that,  $T_{ii}^{(0)} = 0$ 

(2)

 $T_{ii}^{(n+1)} = \max(0, T_{ij}^{(n)} + g_{ij}(x_1, ..., x_{n+1}))$ 

A set  $\{\delta_{ij}$ ,  $1 \le i \le M\}$  of positive thresholds, where  $\delta_{ij}$  is associated with the sequence  $\{T_{ij}^{(n)}, n \ge 1\}$ (3)

The thresholds {\delta\_{ij}, 1\leq i\leq M} are determined a priori based on false alarm and correct decision criteria. The algorithmic decision is then as follows:

Decide that scenario j shifted to scenario k, the first time n, such that  $T_{kj}^{(n)} \ge \delta_{kj}$ , while for each i different than k, the value  $T_{ij}^{(n)}$  is still below its threshold  $\delta_{ij}$ .

The above algorithm is clearly sequential, and numerically very efficient. Furthermore, as proven in [1], it is also asymptotically optimal; that is, there is no other algorithm that can detect a change in model faster than the above algorithm. In addition, if there is lack of confidence in some of the observed data or ambiguity in the models of the battle scenaria, then the algorithm can be easitly modified, to basically maintain its performance characteristics, (see [2]).

#### III. Example

Let the observed data be binary and independent, and let us consider two scenaria, I and II. Let it be known that under scenario I, a bit has value 1 with probability p, and has value 0 with pro auxility 1-p. Let it also be known that under scenario II, a bit has value 1 with probability q, and has value 0 with probability 1-q. Let scenario I be initially present, and let  $x_1, x_2, \ldots$ denote the observed binary sequence. Then, for detecting a change to scenario II, the algorithm easily reduces as follows:

Select some threshold  $\delta$  positive.

Set 
$$T_0 = 0$$
 and  $T_n = \max(0, T_{n-1} + x_n + \frac{\log \frac{1-q}{1-p}}{\log \frac{q(1-p)}{p(1-q)}})$ 

Decide that scenario I shifted to scenario II, the first time n such that  $T_n \ge \delta$ .

Here, we will study the nonasymptotic performance of the algorithm. In particular, given the threshold,  $\delta$ , we will anlayze the false alarm and the power characteristics induced by the algorithm. Let us define.

$$P_{r,\delta}(n)$$
: The probability that the threshold,  $\delta$ , is first crossed, at the time instant,  $n$ , under the condition that the p to q change actually occurred just after the datum  $x_r$ . (4)

The probabilities,  $\{P_{\infty,\delta}(n); n \ge 1\}$ , represent then a false alarm set, while the probabilities,  $\{P_{0,\delta}(n) : n \ge 1\}$ , represent a power set. Indeed, the former probabilities show how probable is the crossing of the threshold, at different time instants, n, when the p to q change never occurs. The latter probabilities correspond, instead, to threshold crossings, at different time instants, n. when the "alarming", parameter-q, process is active, since after time zero. We note that when the threshold,  $\delta$ , is finite, it is eventually crossed with probability one, even when the p to q change never occurs. The false alarm curves,  $\{P_{\infty,\delta}(n); n\geq 1\}$ , and the power curves,  $\{P_{0,\delta}(n), n \ge 1\}$ , can be only computed numerically. For efficiency in such computations, we

apply an approximation. In particular, we approximate the constant, 
$$\frac{\log \frac{1-q}{1-p}}{\log \frac{q(1-p)}{p(1-q)}}$$
 by the ratio of two integers. That is, we write,

$$\frac{\log \frac{1-q}{1-p}}{\log \frac{q(1-p)}{p(1-q)}} \approx -\frac{l}{s}$$
(5)

; where l < s, and l and s are both natural numbers.

Then, we write the algorithm in the following equivalent form, where, without lack in generality, the threshold, t, in the latter is a natural number.

Select the threshold, t. Decide the change has occurred, at the first time, n, such that:  $T'_n \ge t$ , where:

$$T'_0 = 0, T'_n = \max(0, T'_{n-1} + y_n)$$
 (6)

$$y_n = s x_n - l \tag{7}$$

Clearly, the updating step,  $y_n$ , in (7) is an integer. It is equal to -l, when  $x_n = 0$ , and it is equal to s - l > 0, when  $x_n = 1$ . Let us denote by,  $P_{r,t}(n)$ , the probability in (4), as applied to the modified algorithm in (6) and (7). It is then possible to obtain recursive expressions for the computation of the probabilities,  $\{P_{r,t}(n); n \ge 1\}$ , via a Markov chain formulation. The key element in the latter formulation is the probability,  $P_{r,t}(n,j)$ , that  $T'(x^k) < t$ ;  $\forall k < n$  and  $T'(x^n) = j$ , given that the p to q change occurs just after the datum,  $x_r$ . Indeed, we find,

i. If,  $t - 1 \ge s > l + 1$ :

$$P_{r,t}(n,0) = (1-v) \sum_{i=0}^{l} P_{r,t}(n-1, i)$$

$$P_{r,t}(n,j) = (1-v) P_{r,t}(n-1, j+l); 1 \le j < s -l$$
(8)

$$P_{r,t}(n,j) = (1-v) \ P_{r,t}(n-1,\ j+l) + v \ P_{r,t}(n-1,\ j-s+l); \ s-l \le j \le t-l-l$$

$$P_{t,t}(n,j) = v P_{r,t}(n-1, j-s+l); t-1-l < j \le t-1$$

ii. If,  $t - 1 \ge s = l + 1$ :

C

$$P_{r,t}(n,0) = (1-v) \sum_{i=0}^{l} P_{r,t}(n-1, i)$$
(9)

$$P_{r,t}(n,j) = (1-v) P_{r,t}(n-1, j+l) + v P_{r,t}(n-1, j-1); 1 \le j \le t-1-l$$

$$P_{r,t}(n,j) = v P_{r,t}(n-1, j-1); t-1-l < j \le t-1$$

: where.

$$v = \begin{cases} p; & \text{if } n \le r \\ q; & \text{if } n > r \end{cases}$$
 (10)

$$P_{r,t}(n) = v \sum_{j=t-s+l}^{t-1} P_{r,t}(n-1, j) ; t-1 \ge s \ge l+1$$
 (11)

The recursive expressions in (8) and (9) allow for the efficient computation of the probability,  $P_{r,t}(n)$  in (11), for every r, and every n.

The false alarm probabilities,  $\{P_{\infty,t}(n); n \ge 1\}$ , and the power probabilities,  $\{P_{0,t}(n); n \ge 1\}$ , evolve as a special case of the probabilities,  $\{P_{r,t}(n)\}$ , and they can be also computed recursively then. Let us now define the following two quantities.

$$\alpha(p,q,n,t) = \sum_{i=0}^{r} P_{\infty,t}(i)$$
 (12)

$$\beta(p,q,n,t) = \sum_{i=0}^{n} P_{0,t}(i)$$
 (13)

The quantities,  $\alpha(p,q,n,t)$  and  $\beta(p,q,n,t)$ , are the probabilities that the threshold, t, is crossed by the algorithm in (6) and (7), before or at the time instant, n, given that all the data are respectively generated by the p-parameter and the q-parameter process. Thus,  $\alpha(p,q,n,t)$  and  $\beta(p,q,n,t)$  are respectively the false alarm curve and the power curve induced by the algorithm in (6) and (7). Those two curves represent the nonasymptotic performance of the algorithm. We computed those curves for p = .01 and q = .05, and for various choices of the threshold,  $\delta$ . In our computations, we applied the rational approximations represented by expression (5) and the algorithm in (6) and (7), and we used the recursions in (8) ad (9). Our results are shown in Figure 1. From this figure, the powerful nonasymptotic characteristics of the algorithm are evident. For relatively small sample sizes, it simultaneously attains both low false alarm and high power. The choice of the threshold,  $\delta$ , will be determined from the specific false alarm and power requirements, and it will depend on where the emphasis (power versus false alarm) is placed. When the probabilities, p and q, are further apart from each other, than the pair, p = 0.1 and q =.05, then the false alarm and the power curves are drawn further apart, than those in Figure 1. The performance characteristics of the algorithm are then further improved. Studies similar to those in this example can be performed, when dependent binary data structures are instead considered, (see [5]), or when Gaussian data are present.

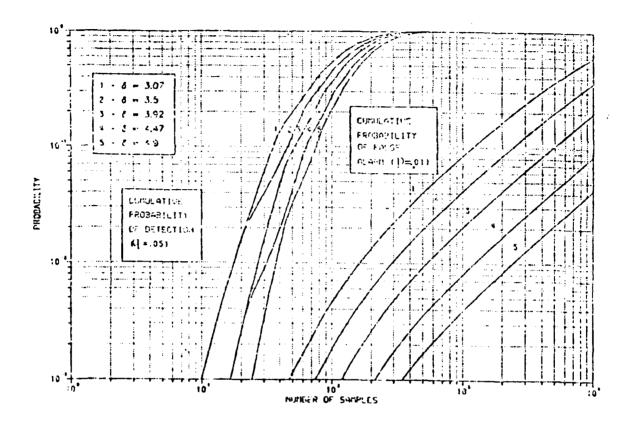


Figure 1

#### IV. CONCLUSIONS

In this paper, we exhibited a simple sequential algorithm which we have previously fully analyzed, (see [1] and [2]), and which can be deployed to detect changes in battle scenaria. The algorithm can be implemented in an automated fasion. At the same time, it would be interesting to see how closely can trained human subjects approximate the algorithmic operations and how fast can they react to the algorithmic alerts. We propose that experiments along those lines be performed by the researchers on human factors, within the DTDM program.

#### REFERENCES

- [1] R. K. Bansal and P. Papantoni-Kazakos, "An Algorithm for Detecting a Change in A Stochastic Process", <u>IEEE Trans. Inform. Th.</u>, IT-32, pp. 227-235, March 1986.
- [2] R. K. Bansal and P. Papantoni-Kazakos, "Outlier Resistant Algorithms for Detecting a Change in Stochastic Process", Univ. of Connecticut, ESE Dept., Technical Report TR-87-1, Jan. 1987. Also, submitted for publication.
- [3] L. Georgiadis and P. Papantoni-Kazakos," A 0.487 Throughput Limited Sensing Algorithm", IEEE Trans. Inform. Th., IT-33, pp. 233-237, March 1987.
- [4] M. Paterakis, L. Georgiadis, and P. Papantoni-Kazakos, "A Full Sensing Window Random-Access Algorithm for Messages with Strict Delay Constraints", Algorithmica, to appear.
- [5] R. K. Bansal, "An Algorithm for Detecting a Change in Stochastic Process", M.S. Thesis, Univ. of Connecticut, 1983.

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